
Factors influencing growth traits in the Nguni cattle stud at Bartlow Combine

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Data collected from Nguni calves born at the Bartlow Combine Breeding Station were used to investigate factors affecting live weights at birth (1960–1991), 205 days (1960–1991), 365 days (1974–1990) and 540 days (1975–1990) of age. Sex, year of birth, sex of calf and age of dam were found to be highly significant (P < 0.001) sources of variation. The effects of interaction between year of birth and sex of calf and the inbreeding of the calf and dam were significant for some traits but their relative effects were small. A suitable model to describe the data for the respective traits which can be used to calculate genetic parameters and breeding values is constructed and discussed.

Data van Ngunikalwers wat by die Bartlow Combine Teelstasie gebore is, is ontleed om faktore wat gewig by geboorte (1960–1991), 205 dae (1960–1991), 365 dae (1974–1990) en 540 dae (1975–1990) beïnvloed, te bepaal. Vader, jaar van geboorte, geslag van kalf en ouderdom van moeder was hoogst betekenisvolle (P < 0.001) bronne van variasie ten opsigte van al die eierskappe. Die effek van interaksie tussen jaar van geboorte en geslag van die kalf en die inteling van kalf en moeder was betekenisvol vir sommige eierskappe maar die relatiewe invloed daarvan was laag. 'n Toepaslike model wat die data van die onderskeie eierskappe beskryf en vir die berekening van genetiese parameters en teelwaardes gebruik kan word, word saamgestel en bespreek.

Keywords: Environmental factors, growth traits, Nguni cattle, non-genetic factors, sire effects.

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Introduction

In the past, indigenous cattle in southern Africa were mainly used as foundation stock for upgrading with Bos taurus bulls which were thought to be superior (Schutte, 1935; Faulkner, 1947; Walker, 1952; Vorster, 1962). These attempts failed in tropical and subtropical regions and attention then became focused on development of various purebred indigenous cattle breeds (Bonsma et al., 1950; Walker, 1964).

The ability of the Nguni to produce and reproduce under harsh environmental conditions (Barnard & Venter, 1983; Scholtz, 1988), their natural immunity against endemic diseases (Spickett & Scholtz, 1985) and suitability as a dam line in terminal crossbreeding (Hofmeyr, 1974; Scholtz, 1988) have generated interest in the breed amongst many cattle farmers. In 1983 the Nguni became a recognized breed under the Livestock Improvement Act (Act 25 of 1977) and the Nguni Cattle Breeders’ Society currently consists of about 120 registered herds with approximately 15 000 registered cattle.

The Bartlow Combine Breeding Station was established in approximately 1954 to accommodate the Nguni stud which originated in 1931. The aim was to investigate the growth, production and reproduction potential of the Nguni cattle as recommended by the Committee which was appointed in 1947 to report on the desirability and means of preserving indigenous livestock (Bonsma et al., 1950). The stud has subsequently played a significant role in the development of the breed and has formed the foundation of a number of the registered herds. The aim of this study was to evaluate factors contributing to the variation in birth weight, 205-, 365- and 540-day weight of Nguni cattle at the Bartlow Combine Breeding Station in order to derive a model needed to estimate genetic and phenotypic parameters for the calculation of breeding values and the development of selection procedures.

Material and Methods

Environment and management

Data were obtained from the Nguni cattle stud kept at the Bartlow Combine Breeding Station. The farm covers an area of 7496 ha and is situated in Zululand, 30 km south of Mkuzi at lat. 32° 03' E and long. 27° 45' S. The altitude varies from 210 to 678 m above sea-level and the topography is hilly to steep.

The climate is subtropical with a mean annual rainfall of approximately 580 mm and four dry months from May to August. Average maximum (minimum) summer and winter temperatures are 29.1°C (17.8°C) and 26.6°C (11.3°C). The vegetation is classified according to Acock's (1975) as veld types 6(a), Zululand Thornveld, and 10, Lowveld. It varies from a sourish mixed open bushveld to a good sweet veld. Acacia spp., Themeda triandra and Panicum spp. are predominant. The area is heavily tick-infested with Boophilus spp., Amblyomma hebraeum, Hyalomma rufipes and Hippolæphalus spp. commonly found.

The stud originated from Nguni cattle purchased and bred since 1931. It was closed from 1957 until 1981 when 20 females, and in 1983 when 18 females and one bull were introduced. In 1960, the breeding animals consisted of 14 bulls and 394 cows and heifers. Assortative mating was practised until 1975. Selection was mainly for type, conformation and...
weight at weaning and two years of age. Milk production of the dam and performance of progeny were important factors in the selection of breeding bulls (Hamburger, 1960–1968). Other criteria included in the selection programme were temperament, hypoplasia, albinism and hollow backs. After 1970, more emphasis was placed on fertility (Hamburger & Swanepoel, 1978; Ramsay & Swanepoel, 1981; Lepen, 1986) as well as growth performance as recorded with the National Performance and Progeny Testing Scheme under Phase A and from 1974, Phase B (Armstrong, 1975; Ramsay, 1981). Breeding bulls were also selected on growth performance as recorded for either Phase C or D (Lepen, 1988).

In 1976, seven breeding lines were formed and were based on the relationship of animals with the original herd sires which were used in 1958 (Hamburger & Swanepoel, 1978). Line breeding was applied in the 1976 breeding season, but subsequently breeding bulls of a particular breeding line were mated to females of other lines (Pretorius, 1980). Single sire mating was practised and sires were replaced as required.

An autumn and spring breeding season of 90 and 65 days, respectively, was used until 1974 after which one season of 90 days per year was retained. Since 1976, it was decreased to 65 days for cows and 45 days for heifers. The latter commenced one month in advance of the cows. More than the required number of replacement heifers were included for mating (Hamburger & Swanepoel, 1978) with final selection on reproduction rate and progeny performance (Lepen, 1988).

The stud was registered with the Nguni Cattle Breeders’ Society in 1986 and it was hence required that breeding animals should comply with the minimum breed standards.

Statistical analysis
The data consisted of records collected from calves at birth (1960–1991), 205 days (1960–1991), 365 days (1974–1990) and 540 days (1975–1990) of age according to the procedures prescribed by the National Beef Cattle Performance and Progeny Testing Scheme (Department of Agriculture and Water Supply, 1988). Except for birth weight (BW), the actual weight of individual calves was corrected for each trait according to the following example:

\[
205\text{-day weight} = BW + 205 \left( \frac{\text{actual WW} - BW}{\text{age in days at weaning}} \right)
\]

where BW = birth weight and WW = weaning weight.

Inbreeding coefficients were calculated for all animals using the algorithm described by Quaas (1976).

The data were edited to exclude the following records:
(i) animals without positive sire identification,
(ii) sires with less than 10 progeny,
(iii) animals born during the period 16 January to 14 August,
(iv) progeny of dams with unknown dates of birth, and
(v) progeny of dams younger than 957 days and older than 6072 days.

Analyses of variance were conducted using least-squares mixed model procedures (Harvey, 1988). Preliminary analyses included sires nested within breeding lines, year of birth, sex of calf, the interaction between year and sex as well as the linear and quadratic regressions of age of dam and the inbreeding of calf and dam on performance.

It was decided to fit the following model:

\[
Y_{ijkl} = \mu + r_i + y_j + s_k + b_1X + b_2X^2 + e_{ijkl}
\]

where

\[
Y_{ijkl} = \text{observation on the } j\text{th calf of the } k\text{th sex, born in the } j\text{th year from sire } i,
\]

\[
\mu = \text{overall mean},
\]

\[
r_i = \text{random effect of the } i\text{th sire with zero mean and variance } \sigma^2_r,
\]

\[
y_j = \text{fixed effect of the } j\text{th year of birth (} j = 1, \ldots, 32),
\]

\[
s_k = \text{fixed effect of the } k\text{th sex (} k = 1, 2),
\]

\[
b_1, b_2 = \text{linear and quadratic regression coefficient of difference between age of dam of the } k\text{th calf (X) and mean age of dams (X) in population},
\]

\[
e_{ijkl} = \text{random error with zero mean and variance } \sigma^2_e.
\]

Results and Discussion
Analysis of variance for the various traits is presented in Table 1. All effects included in the final model were found to be highly significant (\(P < 0.001\)) sources of variation for all the traits under study. Estimates of least squares and standard errors are reported in Table 2.

Breeding lines were found to be a non-significant effect and sires were subsequently included as the only random effect to make provision for genetic differences among calves.

The effect of inbreeding of the calf as a linear regression on birth and 205-day weight was not significant, but was significant (\(P < 0.005\)) for 365- and 540-day weight. Alexander & Bogart (1961) also found inbreeding of the calf not to have a

<table>
<thead>
<tr>
<th>Source</th>
<th>df</th>
<th>MS</th>
<th>df</th>
<th>MS</th>
<th>df</th>
<th>MS</th>
<th>df</th>
<th>MS</th>
</tr>
</thead>
<tbody>
<tr>
<td>Sire</td>
<td>241</td>
<td>72.44</td>
<td>241</td>
<td>1406.94</td>
<td>144</td>
<td>1033.77</td>
<td>122</td>
<td>1219.14</td>
</tr>
<tr>
<td>Year</td>
<td>31</td>
<td>511.01</td>
<td>31</td>
<td>10909.93</td>
<td>16</td>
<td>50383.95</td>
<td>16</td>
<td>50320.00</td>
</tr>
<tr>
<td>Sex</td>
<td>1</td>
<td>10492.28</td>
<td>1</td>
<td>483789.22</td>
<td>1</td>
<td>200866.17</td>
<td>1</td>
<td>92842.42</td>
</tr>
<tr>
<td>Age of dam</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>b1</td>
<td>1</td>
<td>1034.65</td>
<td>1</td>
<td>62959.74</td>
<td>1</td>
<td>22130.54</td>
<td>1</td>
<td>5865.73</td>
</tr>
<tr>
<td>b2</td>
<td>1</td>
<td>1815.76</td>
<td>1</td>
<td>33148.06</td>
<td>1</td>
<td>16409.51</td>
<td>1</td>
<td>9343.43</td>
</tr>
<tr>
<td>Error</td>
<td>12150</td>
<td>11.91</td>
<td>11732</td>
<td>287.22</td>
<td>4881</td>
<td>309.21</td>
<td>3456</td>
<td>494.66</td>
</tr>
</tbody>
</table>

df = degrees of freedom; MS = mean squares; linear (b1) and quadratic (b2) regressions.
For all values: \(P < 0.001\).
significant effect on birth weight. However, significant effects have been reported on birth weight (Swiger et al., 1961) and weaning weight (Koch, 1951; Burgess et al., 1954). Inbreeding of dam was significant ($P < 0.05$) for 205-day weight as a linear regression. Similar effects were reported by Burgess et al. (1954), Koch (1951) and Swiger et al. (1961). The importance of inbreeding of the dam on 205-day weight indicates the effect it has on her milk production during the preweaning growth phase of her calf. The inbreeding coefficients were relatively low with a mean of 0.5% and 1.6% for calves and dams, respectively. In total, 94.3% of the calves had a coefficient of less than 5%. The annual mean and highest coefficient of inbreeding are presented in Figure 1.

It is possible that the coefficients were higher than estimated as it was assumed that the base population was non-inbred. The interaction between year of birth and sex of calf was found to be significant ($P < 0.005$) for birth and 205-day weights. However, the effects of inbreeding and interaction were excluded from the final model since they reduced the residual variance only marginally while increasing the standard errors of the main effects.

Differences between year of birth were highly significant ($P < 0.001$) for all traits under study. This may be expected due to the prevailing extensive farming conditions. Year effects are mainly caused by climate and its influence on the availability and quality of the pasture (Carles & Riley, 1984) and milk production of the dam (Shelby et al., 1955). Year effects will be discussed in a study on environmental trends.

Bull calves were 7.3%, 8.8%, 7.9% and 4.8% heavier than heifers at birth, 205 days, 365 days and 540 days of age, respectively. The significant influence of sex of calf on weight has been reported in various studies (Pahnish et al., 1961;

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**Table 2** Least-squares mean for birth, 205-, 365- and 540-day weight

<table>
<thead>
<tr>
<th>Effect</th>
<th>205-day weight</th>
<th>365-day weight</th>
<th>540-day weight</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Birth weight</td>
<td>Birth weight</td>
<td>Birth weight</td>
</tr>
<tr>
<td>$n$</td>
<td>$n$</td>
<td>$n$</td>
<td>$n$</td>
</tr>
<tr>
<td>$\mu$</td>
<td>12.426</td>
<td>12.008</td>
<td>5.045</td>
</tr>
<tr>
<td>CV (%)</td>
<td>14.55</td>
<td>15.4</td>
<td>12.78</td>
</tr>
<tr>
<td>Sex</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>$\delta$</td>
<td>6.269</td>
<td>6.036</td>
<td>2.161</td>
</tr>
<tr>
<td>$\varphi$</td>
<td>6.157</td>
<td>5.972</td>
<td>2.884</td>
</tr>
<tr>
<td>Age of dam</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>$b_1$</td>
<td>$0.50 \times 10^{-3}$</td>
<td>$3.52 \times 10^{-3}$</td>
<td>$3.93 \times 10^{-3}$</td>
</tr>
<tr>
<td>$\pm 0.05 \times 10^{-3}$</td>
<td>$\pm 0.24 \times 10^{-3}$</td>
<td>$\pm 0.47 \times 10^{-3}$</td>
<td>$\pm 0.73 \times 10^{-3}$</td>
</tr>
<tr>
<td>$b_2$</td>
<td>$-3.1 \times 10^{-5}$</td>
<td>$-13.4 \times 10^{-5}$</td>
<td>$-16.8 \times 10^{-5}$</td>
</tr>
<tr>
<td>$\pm 0.2 \times 10^{-7}$</td>
<td>$\pm 1.3 \times 10^{-7}$</td>
<td>$\pm 2.3 \times 10^{-7}$</td>
<td>$\pm 3.6 \times 10^{-7}$</td>
</tr>
</tbody>
</table>

LSM = least-squares mean; SE = standard error; CV (%) = coefficient of variation.
Andersen & Plum, 1965; Bair et al., 1972; Bailey et al., 1972.

Age of dam fitted as a quadratic regression was found to be highly significant (P < 0.001). Most reports indicate a curvilinear relationship between age of dam and weight of her calves with the highest values from 6 to 10 years for Bos indicus (Venter, 1977). Gregory et al. (1985) also found that age of dam affects birth and weaning weight significantly (P < 0.01), whereas Swanepeol & Heyns (1988) and Bosman & Harwin (1967) found no significant effect of age of dam on weaning weight in the case of Afrikaner cattle. Because some selection had already taken place at weaning and records of the inherent productivity of cows born in later years were not included, post-weaning traits are biased under a sire model with no relationships other than half-sibs.

Conclusions

Year of birth, sex of calf and age of dam were found to be highly significant (P < 0.001) sources of variation in weight in the Nguni stud at Bartlow Combine which must be taken into consideration when selection procedures are developed. Inbreeding of the calf and dam, as well as interaction between year of birth and sex of calf was also highly significant in some of the traits under study. It is not recommended that these effects are included in a model to describe the data as it reduced the residual variance marginally while the standard errors increased. The inconsistency of published results indicates the importance of estimating factors that affect traits within a specific herd and environment.

Acknowledgements

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References

Effect of melatonin implants, flushing and teasing on the reproductive performance of spring-mated Dohne Merino ewes

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The effect of melatonin implants, flushing and teaser rams on the reproductive performance of spring-mated Dohne Merino ewes was investigated. The ewes (n = 534) were randomly allotted to four treatments, namely (1) no treatment (control), (2) melatonin implants (melatonin), (3) flushing and teaser rams (flushing and teasers), and (4) melatonin implants, flushing and teaser rams (combined treatment). Ewes in groups 2 and 4 were implanted with melatonin 32 days before commencement of the mating season (16 November 1990). In groups 3 and 4, ewes were supplemented with 400 g of chocolate male/head/day from three weeks before to the end of the mating period and vasectomized rams were introduced 10 days prior to mating. Conception rate, lambing percentage and fecundity were significantly higher (P < 0.05) in the melatonin treatment group (group 2) (93.5%, 105.4% and 1.13, respectively) than in control ewes (86.6%, 91.0% and 1.05, respectively). A higher (P < 0.05) lambing percentage was recorded in group 3 than in the control group (99.3% vs. 91.0%). When compared with the control treatment (5 December, SE ± 4.5 days), the mean dates of the first recorded oestrus (‘first oestrus’) during mating were 21 November (SE ± 1.6 days), 28 November, (SE ± 3.98 days) and 30 November (SE ± 5.14 days) for groups 2 (P < 0.01), 3 (P < 0.05) and 4 (P < 0.05), respectively. When compared with the control treatment (33 days), the spread of first oestrus was 18, 28 and 27 days for groups 2 (P < 0.01), 3 (P < 0.05), and 4 (P < 0.05), respectively. However, no synergistic effect on reproduction and date of first oestrus above group 2 was found when melatonin-implanted ewes received the supplementation and stimulation with teasers (group 4). In animals treated with melatonin only (group 2) the onset of oestrus was earlier and the reproduction performance was higher than in other treatments (P < 0.05). At current prices melatonin treatment is more cost-effective than flushing and teasing.

Die effek van melatonienimplantering, prikkelvoeding en koggelramme op die reproduksieprestaties van lentegepaarde Dohne Merino-ooie is ondersoek. Ooie (n = 534) is ewekansig in vier behandelingsgroep, naamlik (1) 'n onbehandelde groep (kontrole), (2) groep met melatonin geïmplanteer (melatonin), (3) met prikkelvoeding en koggelramme (prikkelvoeding en koggelramme), en (4) met melatonienimplantings, prikkelvoeding en koggelramme (gkombineerde behandeling), ingedeel. Ooie in groep 2 en 4 is 32 dae voor die aanvang van paardyd (16 November 1990) met melatonin geïmplanteer. Ooie in groep 2 en 4 het 400 g sjokolademielies/kop/dag vanaf drie weke voor, tot die einde van paardyd ontvang en koggelramme is 10 dae voor paring by ooie in die groep geplaas. Besetting, lampersentasie en eerstelikheid was betekenisvol hoër (P < 0.05) by die melatonienbehandeling (groep 2) (93.5%, 105.4% en 1.13, onderskeidelik) as by die kontrole (86.6%, 91.0% en 1.05, onderskeidelik). Ooie in groep 3 het 'n hoër (P < 0.05) lampersentasie as die kontrole (99.3% vs. 91.0%). In vergelyking met die kontrolegroep (5 Desember, SE ± 4.5 dae) was die gemiddelde datum van die eerste aangetekende estrus (‘eerste estrus’) van ooie tydens paardyd 21 November (SE ± 1.6 dae), 28 November (SE ± 3.98) en 30 November (SE ± 5.14) vir groep 2 (P < 0.01), 3 (P < 0.05) en 4 (P < 0.05), onderskeidelik. In vergelyking met die kontrole (33 dae) was die verspreiding van eerste estrus 18, 28 en 27 dae onderskeidelik vir groep 2 (P < 0.01), 3 (P < 0.05) en 4 (P < 0.05). Geen buitensydige verhoging in reproduksie en aanvang van eerste estrus is bo dié van groep 2 gevind toe melatonienbehandelde ooie ook geprikkel en gekoggel was nie (groep 4). Ooie wat slegs met melatonin behandel is (groep 2) se datum van eerste estrus was vervuil en die reproduksieprestaties was betekenisvol bo dié van ander behandelingen verhoog. Teen heersende pryse is melatonienbehandeling meer kost-effikief as die gebruik van prikkelvoeding en koggelramme.

Keywords: Flushing, melatonin, teasing.

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Introduction

The generally accepted concept that natural grazing has a good feeding value during spring and early summer dictates, from an economic point of view, a spring or early summer lambing season for the greater part of the eastern grassveld regions of South Africa (Adler, 1964). Despite the need for supplementary feeding, autumn or early winter lambing is generally recommended owing to a lower incidence of internal parasites during these periods. Reyneke (1969) reported that Dohne Merino ewes mated in autumn produced 27.5% more lambs than ewes mated in spring. A reduced incidence of oestrus in spring and early summer (Shelton & Morrow, 1965; Watson & Radford, 1966) is an important factor which contributes to the low lambing percentage associated with spring mating.